

Solution to Series 4

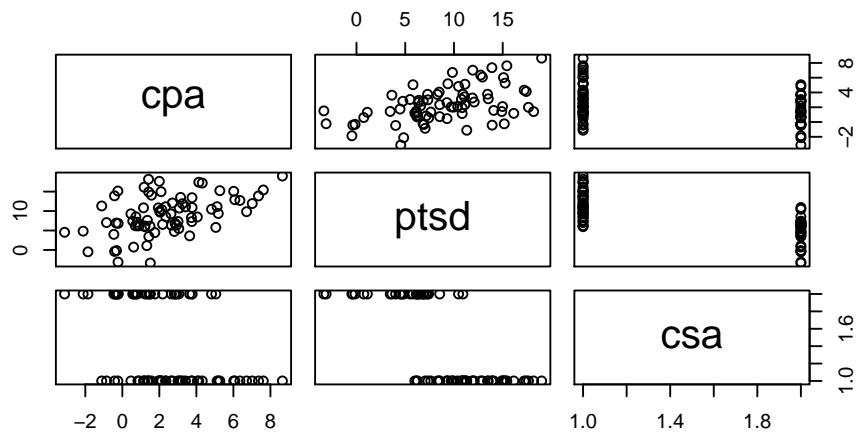
1. a) Read in the data and look at the data, do you see any problems? Make sure that all the variables are in the correct R data type.

```
> sexab <- read.csv("http://stat.ethz.ch/Teaching/Datasets/abuse.csv", header=TRUE)
```

```
> attach(sexab)
```

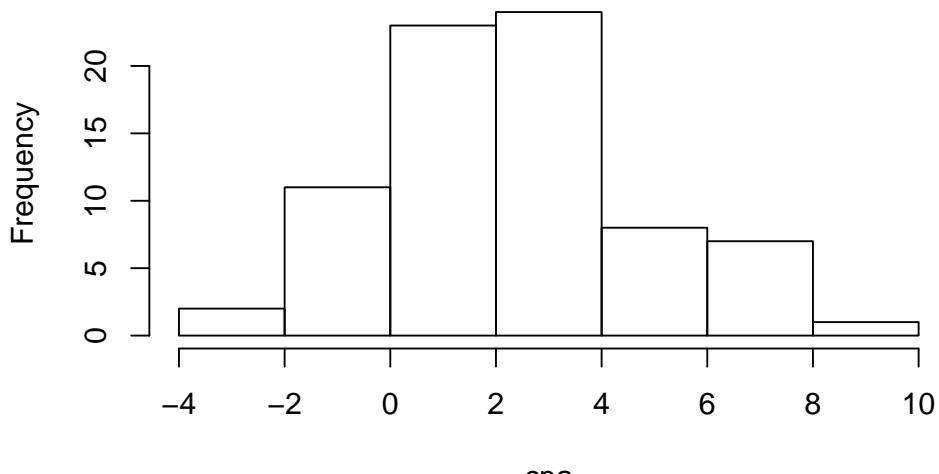
Look at the data:

```
> pairs(sexab)
```

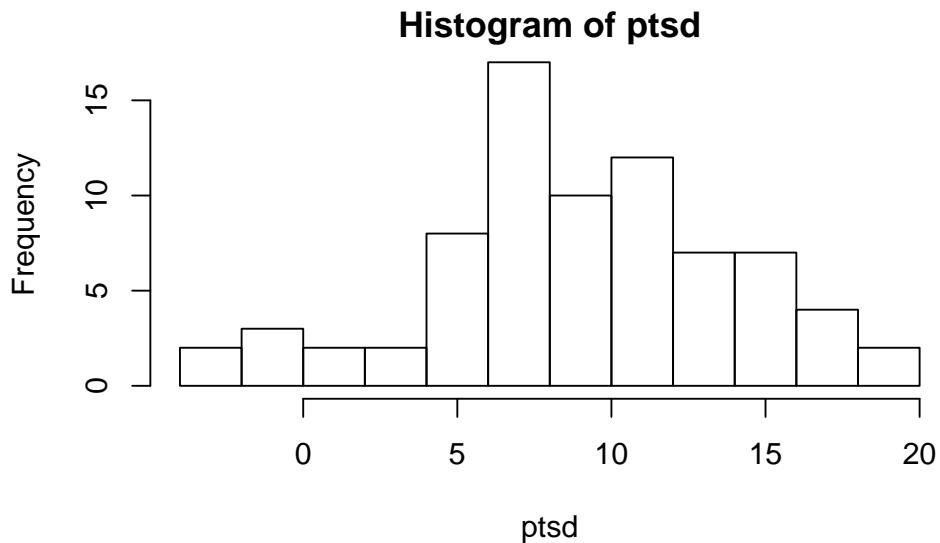


```
> hist(cpa)
```

Histogram of cpa



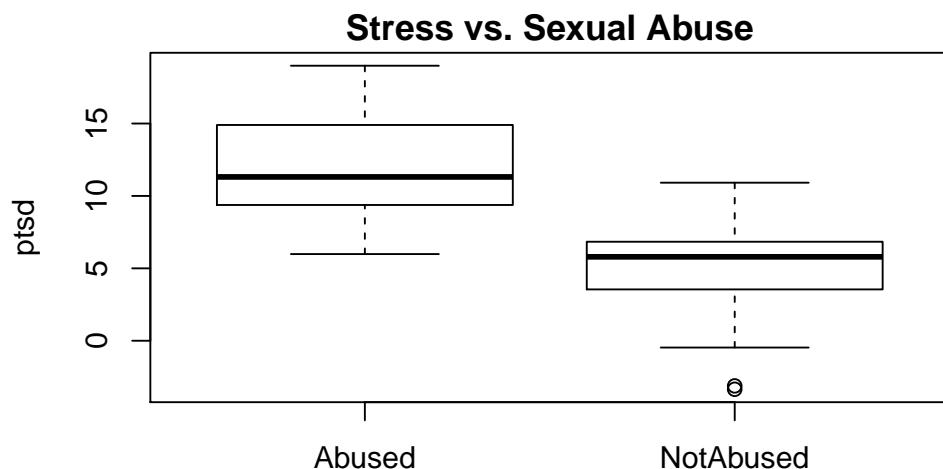
```
> hist(ptsd)
```



No data problems. No transformations necessary.

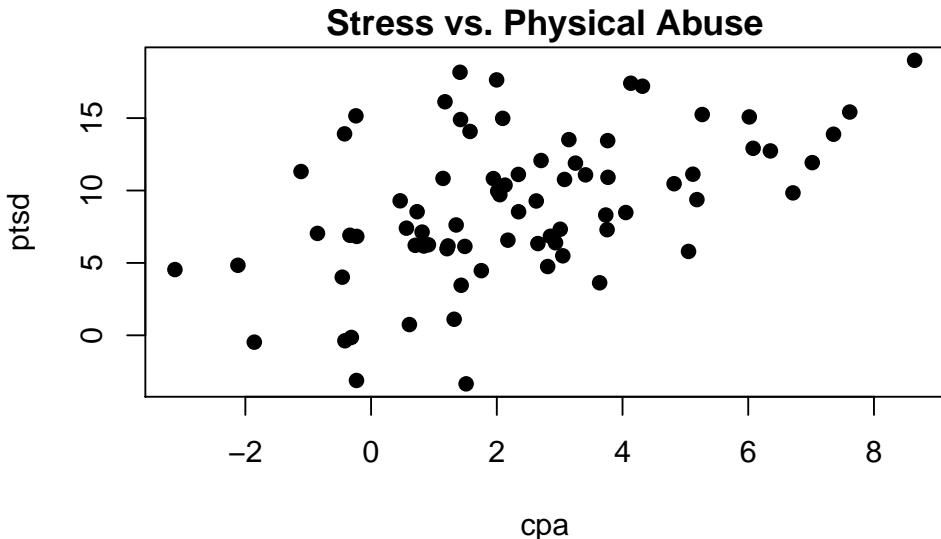
- b) Use scatter plots and box plots to display the variable ptsd in dependence of the variables csa and cpa. Box plot of ptsd vs. csa:

```
> boxplot(ptsд ~ csa, ylab="ptsд", main="Stress vs. Sexual Abuse")
```



Scatter plot of ptsd vs. cpa:

```
> plot(ptsд ~ cpa, ylab="ptsд", main="Stress vs. Physical Abuse", pch=19)
```



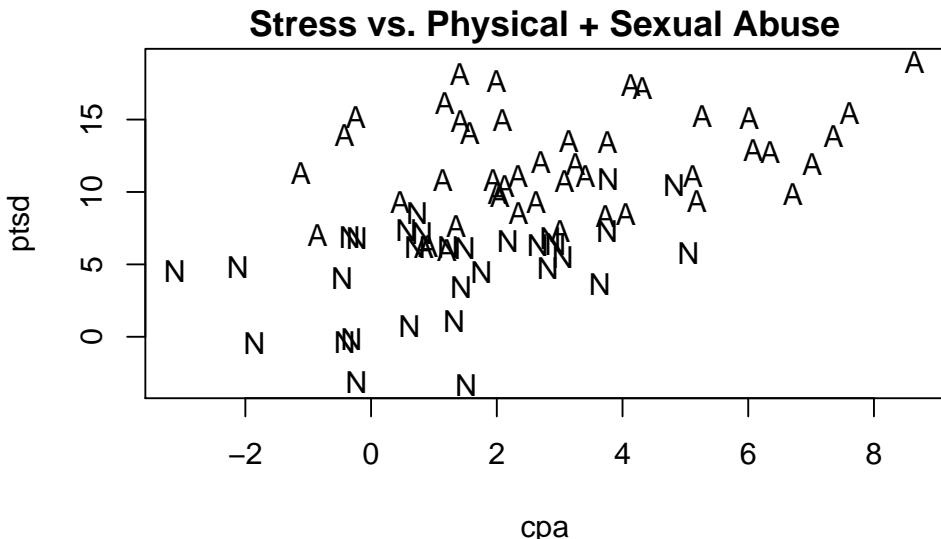
This scatter plot could be misleading. The fact that we plot both groups of woman in one plot could indicate a bigger dependence of ptsd and cpa as there really is.

- c) Make a scatter plot of ptsd against cpa. Use different symbols for abused and non-abused woman.
R-hint:

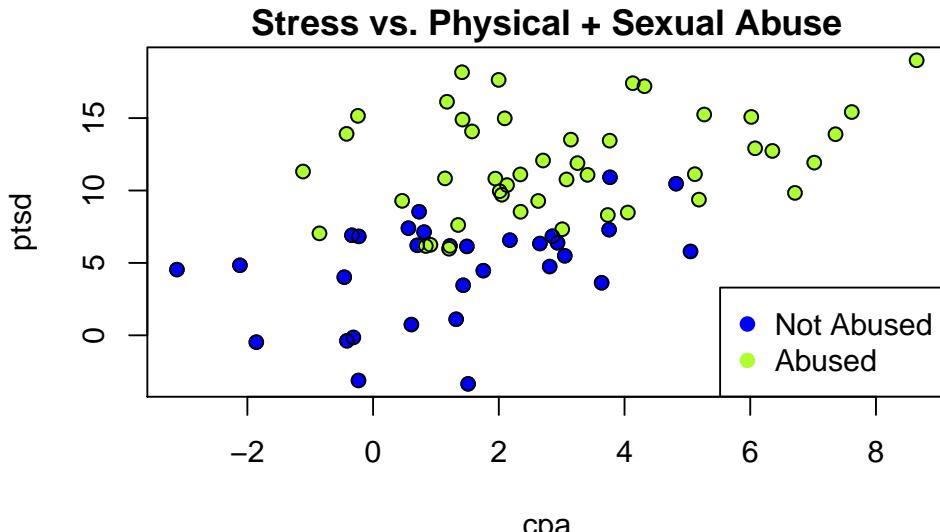
```
plot(cpa, ptsd, type="n")
text(cpa, ptsd, labels=substring(csa,1,1))
```

Scatter plot with different symbols for the different groups.

```
> plot(ptsds ~ cpa, ylab="ptsd", main="Stress vs. Physical + Sexual Abuse", type="n")
> text(cpa, ptsd, labels=substring(csa, 1, 1))
```



```
> plot(ptsds ~ cpa, pch=19, col="blue", main="Stress vs. Physical + Sexual Abuse")
> points(ptsds ~ cpa, pch=19, col="greenyellow", subset=(csa=="Abused"))
> points(ptsds ~ cpa)
> legend("bottomright", legend=c("Not Abused", "Abused"),
       pch=19, col=c("blue", "greenyellow"))
```



From this plots we see that the dependence between ptsd and cpa is not that big. But the two groups differ much concerning the stress-level. We now do a coherent analysis via quantitative methods.

- d) Carry out a test in order to see if sexual abused woman have a higher PTSD-score. Why doesn't this test give you a complete answer? Hint: Look at the scatter plot from part c.).

```
> t.test(ptsд ~ csa, paired=FALSE, var.equal=TRUE)
```

Two Sample t-test

```
data: ptsд by csa
t = 8.9387, df = 74, p-value = 2.172e-13
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 5.630165 8.860273
sample estimates:
 mean in group Abused mean in group NotAbused
 11.941093           4.695874
```

The null-hypothesis gets rejected. This shows us that there is a statistically significant difference in stress-level between the two groups of woman. But what's with the factor physical abuse. We suggest that also the factor physical abuse has a influence on the stress-level. That is, we have to take both variables in to account at the same time. For that we fit a linear regression model.

- e) Fit a regression model to the data with both predictors and their interaction. What do the resulting coefficients mean?

```
> fit.interact <- lm(ptsд ~ cpa * csa, data=sexab)
> summary(fit.interact)
```

Call:

```
lm(formula = ptsд ~ cpa * csa, data = sexab)
```

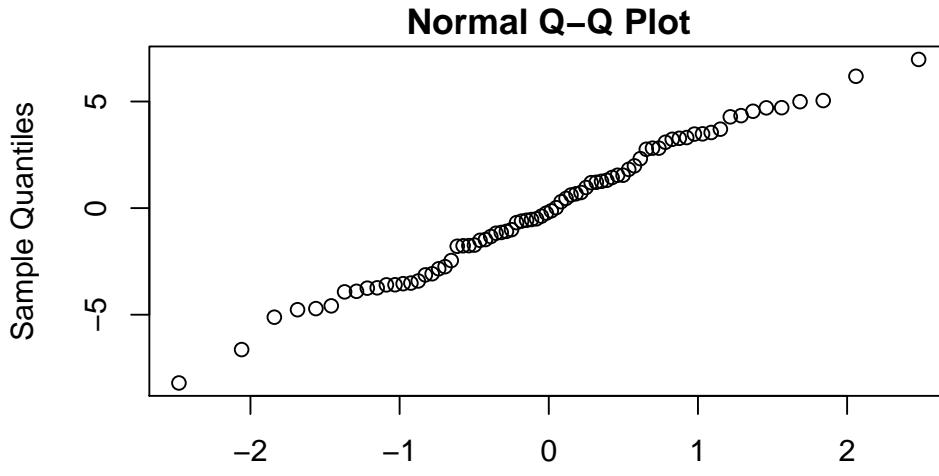
Residuals:

Min	1Q	Median	3Q	Max
-8.1999	-2.5313	-0.1807	2.7744	6.9748

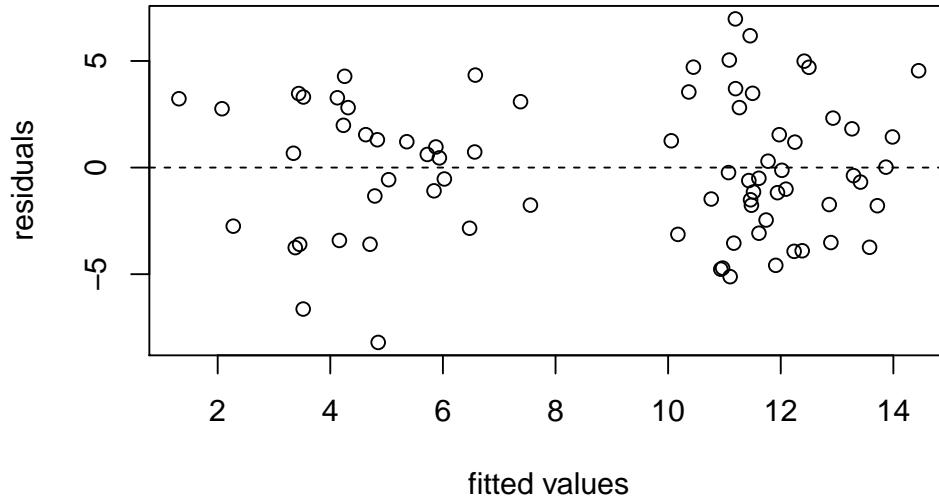
Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	10.5571	0.8063	13.094	< 2e-16 ***
cpa	0.4500	0.2085	2.159	0.0342 *
csaNotAbused	-6.8612	1.0747	-6.384	1.48e-08 ***
cpa:csaNotAbused	0.3140	0.3685	0.852	0.3970

```
---
Signif. codes: 0
> qqnorm(fit.interact$resid)
```



```
> plot(fit.interact$fitted, fit.interact$resid, xlab="fitted values", ylab="residuals")
> abline(h=0, lty=2)
```



2. a) Get an overview of the data and account for possible problems. Which of the variables need to be transformed or not?

Overview over the data:

```
> mortality <- read.csv("http://stat.ethz.ch/Teaching/Datasets/mortality.csv",
  header=TRUE)
> attach(mortality)
> summary(mortality)
```

	City	Mortality
Akron, OH	: 1	Min. : 790.7
Albany-Schenectady-Troy, NY	: 1	1st Qu.: 899.4
Allentown, Bethlehem, PA-NJ	: 1	Median : 946.2
Atlanta, GA	: 1	Mean : 941.2
Baltimore, MD	: 1	3rd Qu.: 984.1

```

Birmingham, AL : 1 Max. :1113.2
(Other) :53

JanTemp JulyTemp RelHum
Min. :12.0 Min. :63.00 Min. :38.00
1st Qu.:27.0 1st Qu.:72.00 1st Qu.:55.50
Median :31.0 Median :74.00 Median :57.00
Mean :33.8 Mean :74.41 Mean :57.75
3rd Qu.:39.5 3rd Qu.:77.00 3rd Qu.:60.00
Max. :67.0 Max. :85.00 Max. :73.00

Rain Educ Dens
Min. :10.00 Min. : 9.00 Min. :1441
1st Qu.:33.50 1st Qu.:10.40 1st Qu.:3138
Median :38.00 Median :11.00 Median :3626
Mean :38.51 Mean :10.97 Mean :3910
3rd Qu.:44.00 3rd Qu.:11.50 3rd Qu.:4566
Max. :65.00 Max. :12.30 Max. :9699

NonWhite WhiteCollar Pop
Min. : 0.80 Min. :33.80 Min. : 124833
1st Qu.: 4.90 1st Qu.:43.40 1st Qu.: 566515
Median : 9.50 Median :45.50 Median : 914427
Mean :11.88 Mean :46.39 Mean :1438037
3rd Qu.:15.70 3rd Qu.:49.90 3rd Qu.:1717201
Max. :38.50 Max. :62.20 Max. :8274961

House Income HC
Min. :2.650 Min. :25782 Min. : 1.00
1st Qu.:3.210 1st Qu.:30004 1st Qu.: 7.00
Median :3.270 Median :32452 Median : 15.00
Mean :3.247 Mean :33247 Mean : 38.47
3rd Qu.:3.360 3rd Qu.:35496 3rd Qu.: 30.50
Max. :3.530 Max. :47966 Max. :648.00

NOx SO2
Min. : 1.00 Min. : 1.00
1st Qu.: 4.00 1st Qu.: 13.00
Median : 9.00 Median : 32.00
Mean : 22.97 Mean : 54.66
3rd Qu.: 24.50 3rd Qu.: 70.00
Max. :319.00 Max. :278.00

```

```

> rownames(mortality) <- mortality$City
> mortality <- mortality[,-1]

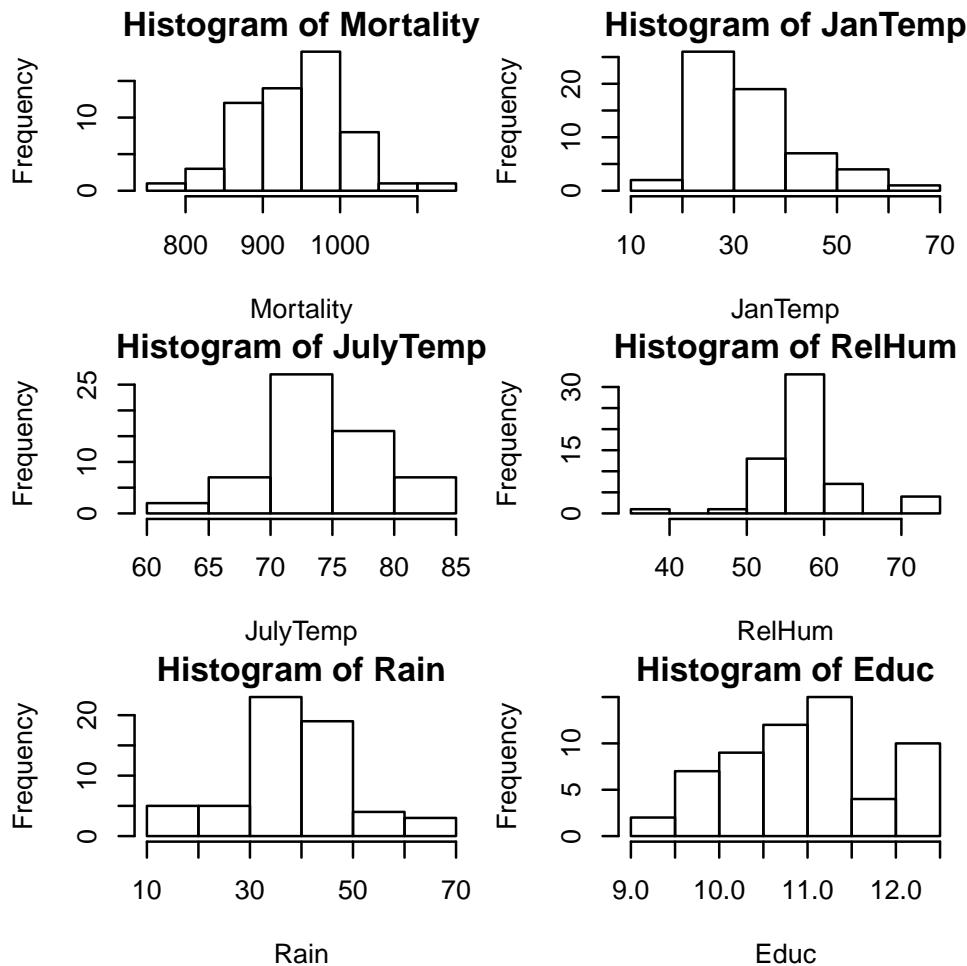
```

We do not see any big data problems. We set city as row names and delete the variable city. Transformationen:

```

> par(mfrow=c(3,2))
> hist(Mortality) ## ok, no transformation
> hist(JanTemp) ## ok, no transformation
> hist(JulyTemp) ## ok, no transformation
> hist(RelHum) ## ok, no transformation
> hist(Rain) ## ok, no transformation
> hist(Educ) ## ok, no transformation

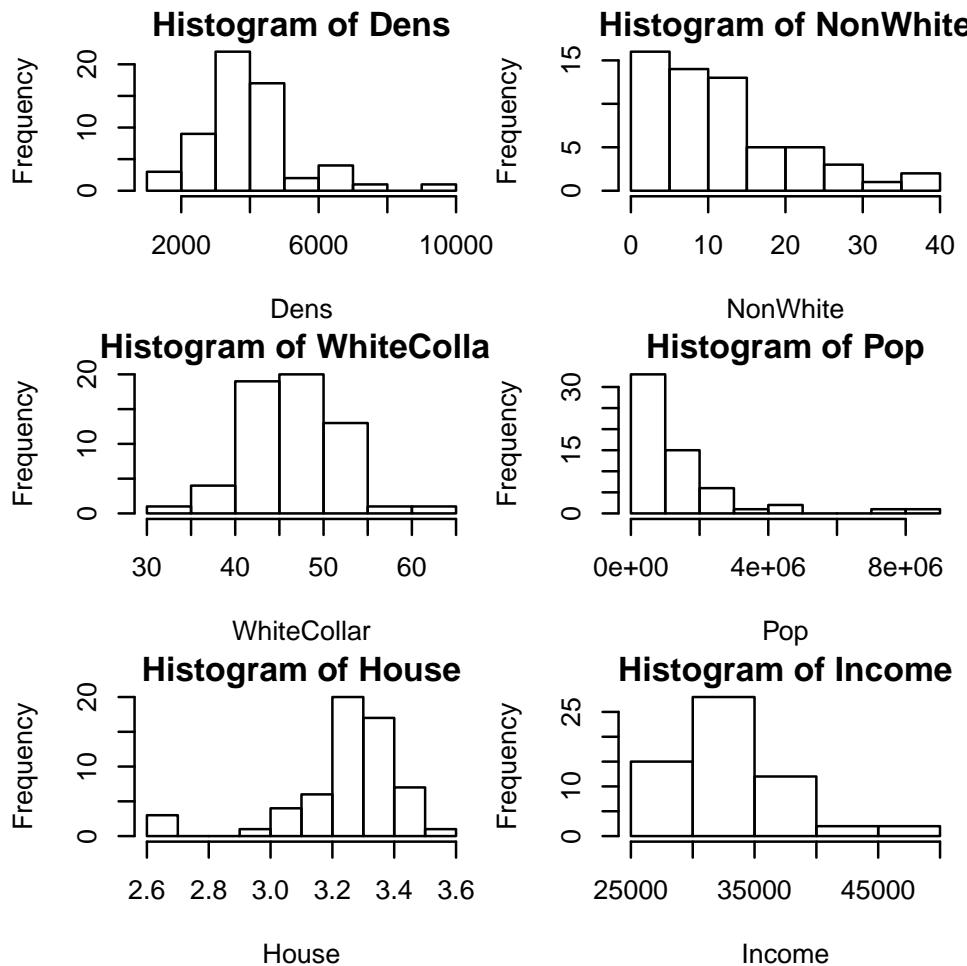
```



```

> par(mfrow=c(3,2))
> hist(Dens)          ## right skewed, log-transformation recomendable
> hist(NonWhite)      ## ratio, arcsin-transformation recomendable
> hist(WhiteCollar)   ## ratio, arcsin-transformation recomendable
> hist(Pop)           ## right skewed, log-transformation recomendable
> hist(House)         ## ok, no transformation
> hist(Income)        ## right skewed, log-transformation recomendable

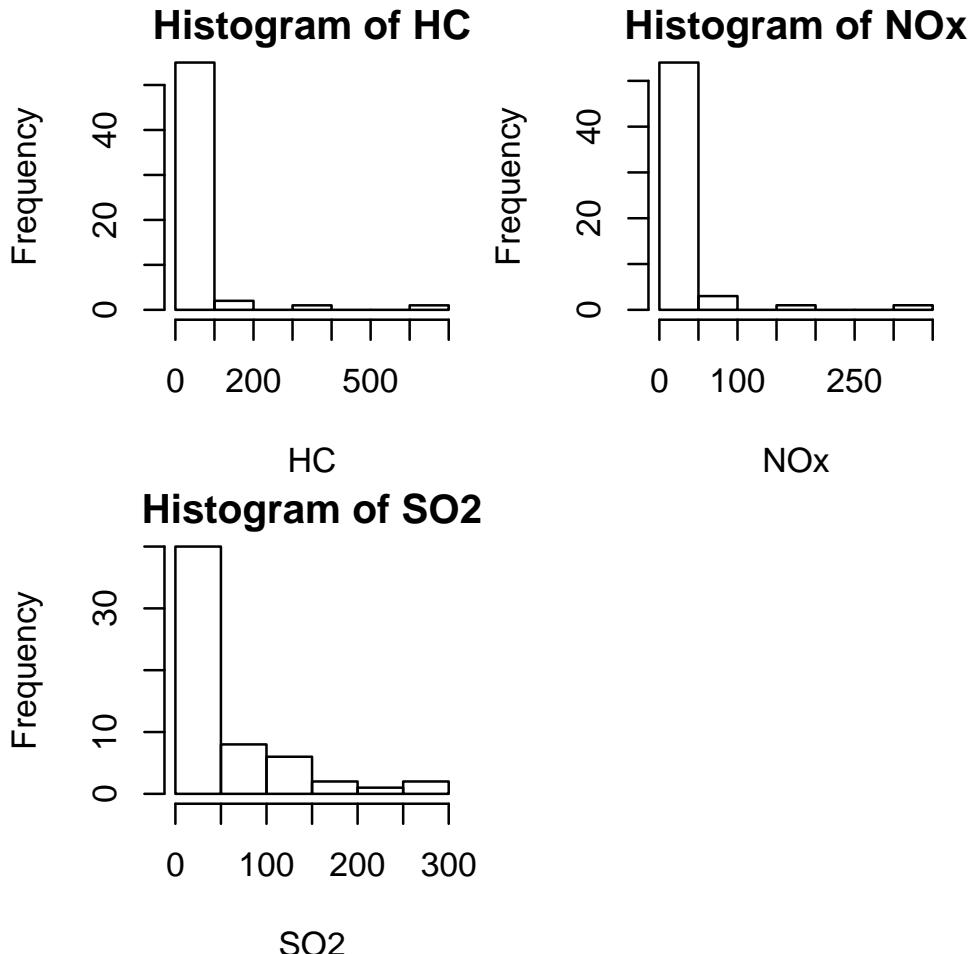
```



```

> par(mfrow=c(2,2))
> hist(HC)          ## strongly right skewed, log-transformation mandatory
> hist(NOx)         ## strongly right skewed, log-transformation mandatory
> hist(SO2)         ## strongly right skewed, log-transformation mandatory

```



```
> detach(mortality)
> mortality$Dens      <- log(mortality$Dens)
> mortality$NonWhite  <- asin(sqrt(mortality$NonWhite/100))
> mortality$WhiteCollar <- asin(sqrt(mortality$WhiteCollar/100))
> mortality$Pop        <- log(mortality$Pop)
> mortality$Income     <- log(mortality$Income)
> mortality$HC          <- log(mortality$HC)
> mortality$NOx         <- log(mortality$NOx)
> mortality$SO2         <- log(mortality$SO2)
> attach(mortality)
```

- b) Carry out a multiple linear regression containing all variables. Does the model fit well? Check the residuals.

Full model:

```
> fit <- lm(Mortality ~ ., data=mortality)
```

```
> summary(fit)
```

Call:

```
lm(formula = Mortality ~ ., data = mortality)
```

Residuals:

Min	1Q	Median	3Q	Max
-65.08	-25.23	-2.67	23.08	75.70

Coefficients:

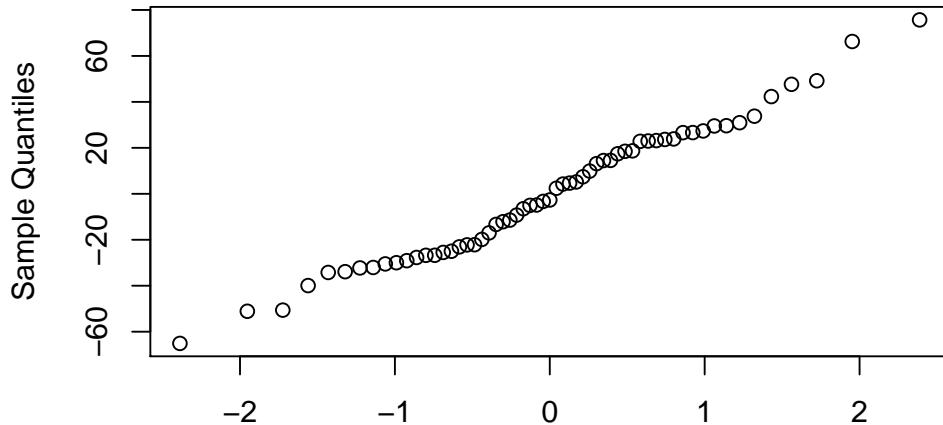
	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	1496.4915	572.7205	2.613	0.01224 *

```

JanTemp      -2.4479    0.8808   -2.779   0.00798  **
JulyTemp     -1.9350    2.0329   -0.952   0.34638
RelHum       0.1065    1.0614    0.100    0.92052
Rain         1.7727    0.5748    3.084    0.00352  **
Educ        -13.3849   8.7561   -1.529    0.13351
Dens         11.9490   16.1836   0.738    0.46423
NonWhite    326.6757  62.9092   5.193 5.09e-06 ***
WhiteCollar -146.3477 112.5510  -1.300    0.20028
Pop          4.8037    7.7245   0.622    0.53723
House        -43.2697  38.9460  -1.111    0.27260
Income       -27.3906  47.8041  -0.573    0.56958
HC           -21.1925  15.1050  -1.403    0.16763
NOx          35.7323   14.3143   2.496    0.01637  *
SO2          -5.3995   7.4040  -0.729    0.46970
---
Signif. codes:  0
> qqnorm(fit$resid)

```

Normal Q-Q Plot

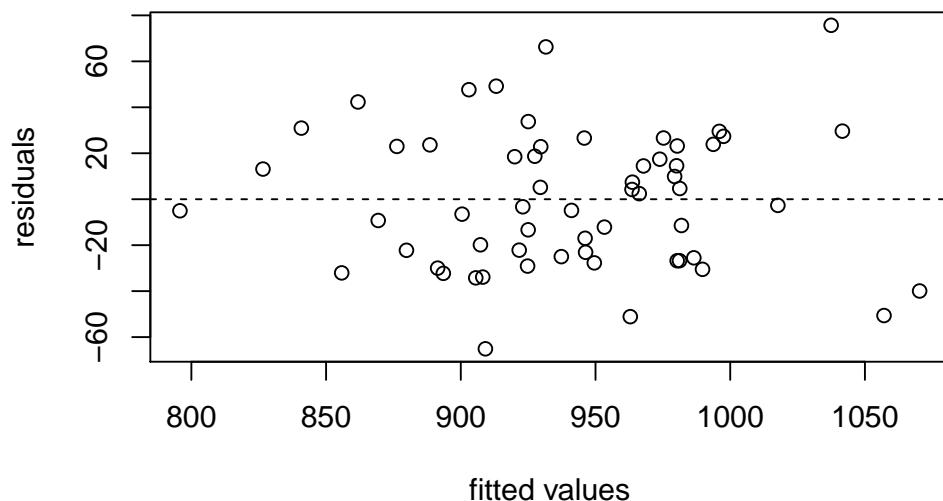


Theoretical Quantiles

```

> plot(fit$fitted, fit$resid, xlab="fitted values", ylab="residuals")
> abline(h=0, lty=2)

```



This model fits quite well, i.e. the model assumptions are fulfilled. We do not see any violation of the model assumptions.

- c) Now take all the non-significant variables out of the model and compute the regression again. Compare your results to the one from part b.).

Now just use the significant variables:

```
> fit.sign <- lm(Mortality ~ JanTemp + Rain + NonWhite + NOx, data=mortality)
> summary(fit.sign)
```

Call:

```
lm(formula = Mortality ~ JanTemp + Rain + NonWhite + NOx, data = mortality)
```

Residuals:

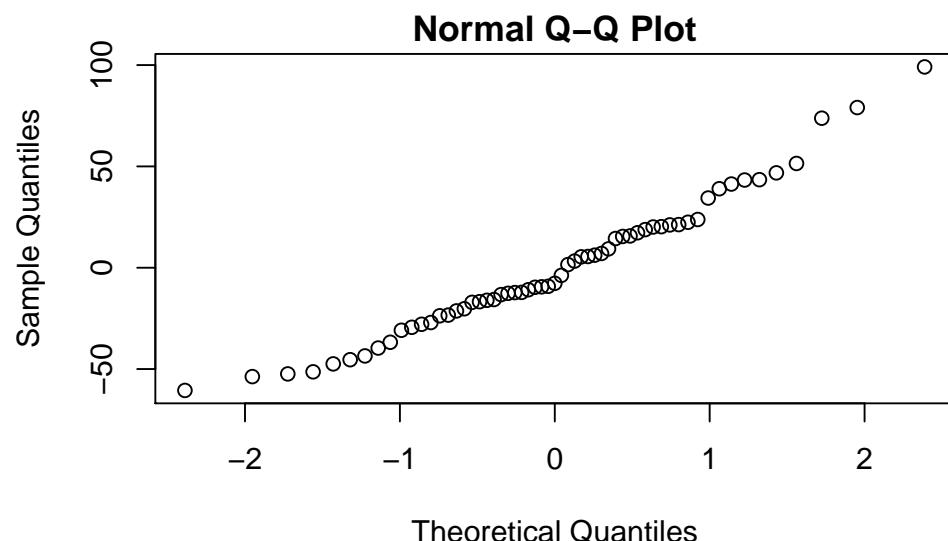
Min	1Q	Median	3Q	Max
-60.537	-22.328	-7.677	20.186	99.117

Coefficients:

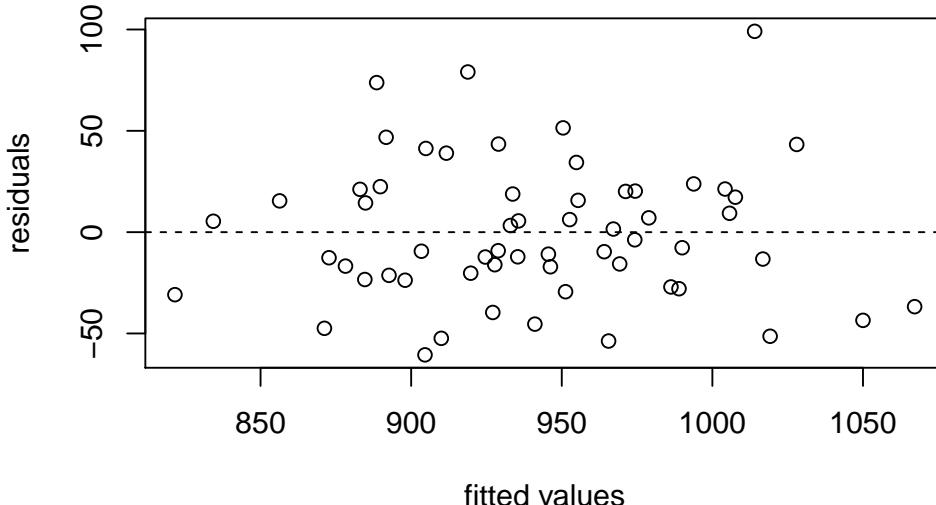
	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	788.6724	25.8034	30.565	< 2e-16 ***
JanTemp	-2.4277	0.5166	-4.699	1.84e-05 ***
Rain	2.4648	0.4692	5.254	2.59e-06 ***
NonWhite	277.1610	40.9045	6.776	9.53e-09 ***
NOx	20.6490	4.5502	4.538	3.21e-05 ***

Signif. codes: 0

```
> qqnorm(fit.sign$resid)
```



```
> plot(fit.sign$fitted, fit.sign$resid, xlab="fitted values", ylab="residuals")
> abline(h=0, lty=2)
```



Now all the variables are highly significant. The error variance is slightly bigger, R-squared and also adjusted R-squared are smaller compared to the full model. On the other hand is the p-value of the F-test bigger now.

Even though leaving away all of the non-significant variables worked quite well here, one should not do that. A better strategy would be to delete the non-significant variables step by step, always deleting the one with the biggest p-value.

- d) Start with the full multiple linear model. Remove now step by step the variable with the biggest p-value as long as it is over 0.05. Compare the result to the one from c.). R-hint: Use the R-function `update()`.

Step by step strategy: Use the function `update()`.

```
> fit.reduc <- fit
> fit.reduc <- update(fit.reduc, ~.-RelHum); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-Income); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-Pop); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-Dens); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-SO2); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-JulyTemp); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-HC); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-House); summary(fit.reduc)
> fit.reduc <- update(fit.reduc, ~.-WhiteCollar); summary(fit.reduc)

Call:
lm(formula = Mortality ~ JanTemp + Rain + Educ + NonWhite + NOx,
    data = mortality)
```

Residuals:

Min	1Q	Median	3Q	Max
-83.471	-23.987	4.444	19.880	85.943

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	992.2069	79.6994	12.449	< 2e-16 ***
JanTemp	-2.1304	0.5017	-4.246	8.80e-05 ***
Rain	1.8122	0.5066	3.577	0.000752 ***
Educ	-16.4207	6.1202	-2.683	0.009710 **
NonWhite	268.2564	38.8832	6.899	6.56e-09 ***
NOx	18.3230	4.3960	4.168	0.000114 ***

Signif. codes:	0			

Now we stop because all of the remaining variables are significant. The error-variance, R-squared and p-value of the F-test look better then in the model from part c.). Also the residuals are looking good.

- e) Again starting from the full model, carry out partial F-tests, in order to answer the question if
- all meteo-variables
 - all air pollution-variables and
 - all demographic-variables

can be removed from the model. Use the R-function `anova()`.

Fitting the model without the meteo-variables:

```
> fit.ohne.meteo <- lm(Mortality ~ .-JanTemp-JulyTemp-RelHum-Rain, data=mortality)
> anova(fit, fit.ohne.meteo)
```

Analysis of Variance Table

```
Model 1: Mortality ~ JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
SO2
```

```
Model 2: Mortality ~ (JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
SO2) - JanTemp - JulyTemp - RelHum - Rain
```

Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	44	51543			
2	48	71705	-4	-20162	4.3027 0.005037 **

Signif. codes: 0

With the function `anova()` one carries out a F-test in order to compare the two models. This test is significant, i.e. the null-hypothesis gets rejected. That is, the bigger model, the one with the meteo-variables, is better. So we can not leave away the meteo-variables.

Fitting the model without the air pollution-variables:

```
> fit.ohne.luft <- lm(Mortality ~ .-HC-NOx-SO2, data=mortality)
> anova(fit, fit.ohne.luft)
```

Analysis of Variance Table

```
Model 1: Mortality ~ JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
SO2
```

```
Model 2: Mortality ~ (JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
SO2) - HC - NOx - SO2
```

Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	44	51543			
2	47	61244	-3	-9700.8	2.7604 0.0533 .

Signif. codes: 0

The partial F-test is not significant. Hence we can take the air pollution-variables out of the model.

Fitting the model without the demographic-variables:

```
> fit.ohne.demografie <- lm(Mortality ~ .-Educ-Dens-NonWhite-WhiteCollar-Pop-House
-Income, data=mortality)
> anova(fit, fit.ohne.demografie)
```

Analysis of Variance Table

```
Model 1: Mortality ~ JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
```

```
S02
Model 2: Mortality ~ (JanTemp + JulyTemp + RelHum + Rain + Educ + Dens +
NonWhite + WhiteCollar + Pop + House + Income + HC + NOx +
S02) - Educ - Dens - NonWhite - WhiteCollar - Pop - House -
Income
Res.Df      RSS Df Sum of Sq      F     Pr(>F)
1        44  51543
2        51 101406 -7   -49863 6.0808 5.369e-05 ***
---
Signif. codes:  0
```

The p-value of the test is very small, that is we can not leave away the demographic-variables.